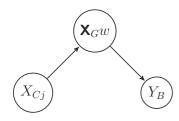




Introduction

Central idea

The role of GE data may be explored by a linear combination of genes (here the columns of X_G), say $X_G w$, so that it maximises simultaneously its correlation with the FF j and with the BA.



The construct X_{GW} is referred to as the gene signature.

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Principal Bicorrelation Analysis

maximisation problem

РВА

A note on the two correlations involved in $R_i^2(w)$:

- Cor $\{Y_B, \mathbf{X}_G w\}^2$: Both Y_B and $\mathbf{X}_G w$ are continuous variables;
- ullet Cor $\{X_{Cj}, \mathbf{X}_G w\}^2$: The FF X_{Cj} is a binary indicator and $\mathbf{X}_G w$ is a continuous variable; hence, the Pearson correlation perhaps is not the best choice (see next slide)

A potential problem: it is easier to find a w to make Cor $\{X_{C_i}, \mathbf{X}_G w\}^2$ large, than it is to make $\text{Cor}\{Y_B, \mathbf{X}_G w\}^2$ large.

Our solution: assign weights $0 < \tau < 1$ and $1 - \tau$ to the two correlations (with τ small).

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Principal Bicorrelation Analysis

maximisation problem

РВА

In particular we are looking for the vector \boldsymbol{w} that maximises the weighted sum of squared correlations (bicorrelation):

$$\begin{split} R_j^2\left(w\right) &= \tau \mathsf{Cor}\left\{X_{Cj}, \mathbf{X}_G w\right\}^2 + (1-\tau) \mathsf{Cor}\left\{Y_B, \mathbf{X}_G w\right\}^2, \\ &= \tau w^\top \mathbf{X}_G^\top X_{Cj} X_{Cj}^\top X_G w + (1-\tau) w^\top \mathbf{X}_G^\top Y_B Y_B^\top \mathbf{X}_G w \end{split}$$

where $0 < \tau < 1$ is a user-defined weight.

The maximisation is subject to the constraints:

- ullet $\|w\|_2=1$, if the desired solution need not to be sparse (i.e. the dense solution)
- $||w||_2 = 1$ and $||w||_1 \le k$ if a sparse solution is desired (k = tuning parameter).

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Principal Bicorrelation Analysis

maximisation problem

РВА

Since the FF X_{Cj} is a binary indicator, the Pearson correlation $\operatorname{Cor}\left\{X_{Cj},\mathbf{X}_{Gw}\right\}^2$ may not be the most approriate measure for association.

We have also evaluated the Ranked Pointwise Biserial coefficient (a mean rank difference),

$$\rho_{rpb} = \frac{2}{n} \left(\bar{r}_1 - \bar{r}_0 \right),\,$$

where $\bar{r}_p=\frac{1}{n_p}\sum_{i:y_i=p}R\{x_i\},\ n_p=\#\{y_i=p\},\ p=0\ {\rm or}\ 1$ and $R\{\cdot\}$ is the rank operator.

But we will show only results for the first one for better interpretation.

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Principal Bicorrelation Analysis

maximisation problem

This w may be obtained from the SVD of the $2 \times n_G$ matrix

$$M := \begin{bmatrix} \tau \mathrm{Cor} \left\{ X_{Cj}, \mathbf{X}_G \right\} \\ (1 - \tau) \mathrm{Cor} \left\{ Y_B, \mathbf{X}_G \right\} \end{bmatrix} \propto \begin{bmatrix} \tau X_{Cj}^\top \mathbf{X}_G \\ (1 - \tau) Y_B^\top \mathbf{X}_G \end{bmatrix} \mathbf{a} \,,$$

- \bullet The dense solution: w is equal to the first right singular vector of M;
- ullet The sparse solution: w is calculated from M by applying the methods described in D. Witten et al. (2009) for sparse PCA; an iterative algorithm with $u^{T}Xv$ as objective function (plus constraints).

^ain the Pearson correlation case, see E. Bair et al. (2004) for details NCS conference, Potsdam



PBA: interesting Plots 1

Description

РВА

Plot of $\operatorname{Cor}\{Y_B, \mathbf{X}_G\hat{w}_1\}^2$ vs $\operatorname{Cor}\{X_{Cj}, \mathbf{X}_G\hat{w}_1\}^2$ for all FFs can be used to select FFs for which $\mathbf{X}_G\hat{w}_1$ is:

- ullet highly correlated with both the BA and FF j
- extremely low correlated with both the BA and FF j, but for which correlation along the direct path FF-BA

Note that in this plot each dot is a FF.

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Principal Bicorrelation Analysis maximisation problem Note 1 all data matrices have been divided by their own first singular value to balance the contribution of РВА each source of data, as in Multiple Factor Analysis. **Note 2** the vector w that maximises $R_i^2(w)$ is denoted Note 3 just like in a PCA, more than one gene signature can be obtained: *i.e.* $\hat{w}_1, \hat{w}_2, \ldots$ September 25th, 2012

