

How to set historical control limits using the R package predint

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Introduction

Aim of the talk

- ▶ Validation of the current control (CC) using historical control limits (HCL)
- ▶ computed based on historical control data (HCD)

OECD Test Guidelines:

- ▶ ... concurrent negative controls should ideally be within the [historical] 95% control limits...

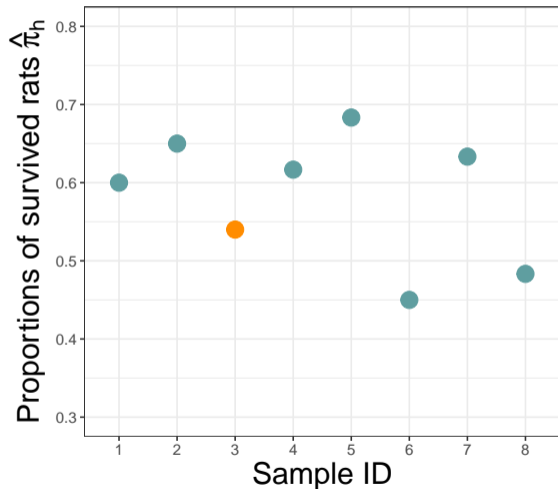
EFSA Scientific opinion (draft)

- ▶ ... a statistical comparison of HCD and CC is performed through calculation of a prediction interval (PI) derived from HCD.

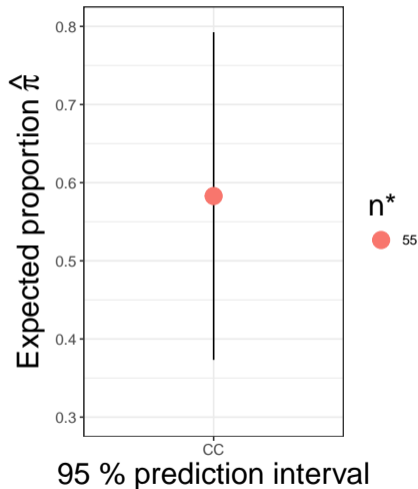
¹ **OECD 471:** Bacterial Reverse Mutation Test; **OECD 473:** In Vitro Mammalian Chromosomal Aberration Test; **OECD 490:** In Vitro Mammalian Cell Gene Mutation Tests Using the Thymidine Kinase Gene; **OECD 2016:** Overview of the set of OECD genetic toxicology test guidelines and updates performed in 2014–2015. **EFSA 2024:** Draft Scientific Opinion on the use and reporting of historical control data for regulatory studies. Public Consultation PC-0856.

Historical control limits (HCL)

Historical control data



Current control (CC)



¹Carlus et al. 2013: Historical control data of neoplastic lesions in the Wistar Hannover Rat among eight 2-year carcinogenicity studies, Experimental and Toxicologic Pathology, 65(3), 243-253.

Hierarchical design

- ▶ n_h experimental units nested within each of the $h = 1, \dots, H$ control groups
- ▶ Systematic between-study variation
- ▶ Different numbers of experimental units per control group possible ($n_h \neq n_{h'}$)

Assumption

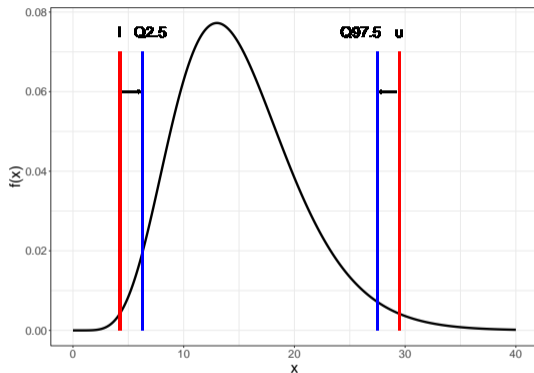
- ▶ HCD and CC derive from the same data generating process

Prediction interval

- ▶ $P(l(\mathbf{y}) \leq \mathbf{y}^* \leq u(\mathbf{y})) = 1 - \alpha$
- ▶ HCD: \mathbf{y}
- ▶ Current control: \mathbf{y}^*

Equal tail probabilities

- ▶ $P(l(\mathbf{y}) \leq \mathbf{y}^*) = 1 - \alpha/2$
- ▶ $P(\mathbf{y}^* \leq u(\mathbf{y})) = 1 - \alpha/2$



Aim

- ▶ Provide methodology for the calculation of prediction intervals for different
 - ▶ experimental designs
 - ▶ model classes
 - ▶ scales

- ▶ Ensure for equal tail probabilities

Bootstrap calibration

Normal approximation

$$\frac{\hat{y}^* - Y^*}{\sqrt{\widehat{\text{var}}(\hat{y}^* - Y^*)}} = \frac{\hat{y}^* - Y^*}{\sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)}} \underset{\text{appr.}}{\sim} N(0, 1)$$

Prediction standard error

$$\widehat{\text{se}}(\hat{y}^* - Y^*) = \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)}$$

Asymptotic prediction interval

$$[l, u] = \hat{y}^* \pm q_{1-\alpha/2} \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)}$$

¹Nelson 1982: Applied life data analysis. Wiley and Sons.

Standardized framework

- ▶ Works for several distributions

Adjust the prediction interval for

- ▶ skewed distributions
- ▶ hierarchical design of HCD
- ▶ simultaneous prediction

Prediction standard error

$$\widehat{se}(\hat{y}^* - Y^*) = \sqrt{\widehat{var}(\hat{y}^*) + \widehat{var}(Y^*)}$$

Pointwise prediction interval

$$[l, u] = \hat{y}^* \pm q_{1-\alpha/2} \sqrt{\widehat{var}(\hat{y}^*) + \widehat{var}(Y^*)}$$

Bootstrap calibration

- ▶ Find substitutes q_l and q_u for $q_{1-\alpha/2}$, that
 - ▶ account for possible skewness
 - ▶ enable simultaneous prediction

$$\left[l = \hat{y}^* - q_l \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)}, u = \hat{y}^* + q_u \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)} \right]$$

The R package "predint"

```
# Install the package from CRAN  
install.packages("predint")  
  
# Install developmental version  
devtools::install_github("MaxMenssen/predint")  
  
# Load the package to current R session  
library(predint)
```

Prediction intervals currently available for

- ▶ Continuous measurements
 - ▶ Linear random effect models
 - ▶ `lmer_pi_futmat()`

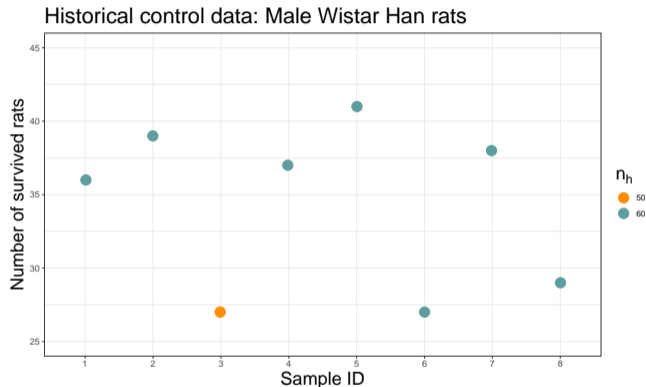
- ▶ Overdispersed binomial data
 - ▶ quasi-binomial: `quasi_bin_pi()`
 - ▶ beta-binomial: `beta_bin_pi()`

- ▶ Overdispersed count data
 - ▶ quasi-Poisson: `quasi_pois_pi()`
 - ▶ negative-binomial: `neg_bin_pi()`

¹ **Menssen and Schaarschmidt 2022:** Prediction intervals for all of M future observations based on linear random effects models, *Statistica Neerlandica*, 76(3), 283-308. **Menssen et al. 2024:** Prediction intervals for overdispersed Poisson data and their application in medical and pre-clinical quality control arXiv:2404.05282 (under review). **Menssen and Rathjens 2024:** Prediction intervals for overdispersed binomial endpoints and their application to historical control data arXiv:2407.13296 (under review).

Binomial data: Survival incidence

##	surv	dead
## 1	36	24
## 2	39	21
## 3	27	23
## 4	37	23
## 5	41	19
## 6	27	33
## 7	38	22
## 8	29	31



¹Carlus et al. 2013: Historical control data of neoplastic lesions in the Wistar Hannover Rat among eight 2-year carcinogenicity studies, Experimental and Toxicologic Pathology, 65(3), 243-253.

Aim

- ▶ Predict the number of rats that survived the two year period
- ▶ Current control group: $n^* = 55$ rats

Beta-binomial assumption

- ▶ $\text{var}(Y_h) = \phi_h n_h \pi(1 - \pi)$ with $\phi_h = 1 + (n_h - 1)\rho$
- ▶ $\text{var}(Y^*) = \phi^* n^* \pi(1 - \pi)$ with $\phi^* = 1 + (n^* - 1)\rho$

- ▶ π binomial proportion
- ▶ ρ intra-class correlation
 - ▶ Overdispersion $\rho > 0$

Prediction standard error

- ▶ $\widehat{se}(\hat{y}^* - Y^*) = \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)} = \sqrt{\frac{\hat{\phi}^* n^{*2} \hat{\pi}(1 - \hat{\pi})}{\sum_h n_h} + \hat{\phi}^* n^* \hat{\pi}(1 - \hat{\pi})}$

¹Messen and Schaarschmidt 2019: Prediction intervals for overdispersed binomial data with application to historical controls. *Statistics in Medicine*, 38(14), 2652-2663.

BS-calibrated interval

$$\left[l = n^* \hat{\pi} - q_l \sqrt{\frac{\hat{\phi}^* n^{*2} \hat{\pi} (1 - \hat{\pi})}{\sum_h n_h} + \hat{\phi}^* n^* \hat{\pi} (1 - \hat{\pi})}, \right.$$
$$\left. u = n^* \hat{\pi} + q_u \sqrt{\frac{\hat{\phi}^* n^{*2} \hat{\pi} (1 - \hat{\pi})}{\sum_h n_h} + \hat{\phi}^* n^* \hat{\pi} (1 - \hat{\pi})} \right]$$

¹Messen and Rathjens 2024: Prediction intervals for overdispersed binomial endpoints and their application to historical control data arXiv:2404.05282 (under review).

```
bb_predint <- beta_bin_pi(histdat = f24a,  
                           newsize=55)
```

```
bb_predint$rho # 0.01169135  
# phi_star = 1 + 54 * 0.011 = 1.63
```

```
summary(bb_predint)
```

```
## Pointwise 95 % prediction interval for one future observation
```

```
##  
##      lower      upper newsize y_star_hat      ql      qu      pred_se  
## 1 20.53592 42.81963      55  32.06383 2.039219 1.902637 5.653101
```

```
##  
## Bootstrap calibration was done for each prediction limit separately  
## using a modified version of Menssen and Schaarschmidt 2022
```

Prediction interval on observation scale (y^*)

▶ $[l, u] = [20.53, 43.59]$

Prediction interval on proportion scale ($\hat{\pi}^*$)

▶ $[l/n^*, u/n^*] = [20.53/55, 43.59/55] = [0.37, 0.79]$

Discussion

Historical control limits

- ▶ HCD based validation of CC is mandatory
- ▶ Literature on prediction intervals for hierarchical designs is relatively scarce
- ▶ More research is needed!

R package predint

- ▶ Prediction intervals for three important scales
- ▶ Coverage probabilities close to the nominal level
- ▶ Provides methodology for self-implementation

¹ **Menssen and Schaarschmidt 2019:** Prediction intervals for overdispersed binomial data with application to historical controls. *Statistics in Medicine*, 38(14), 2652-2663. **Menssen and Schaarschmidt 2022:** Prediction intervals for all of M future observations based on linear random effects models, *Statistica Neerlandica*, 76(3), 283-308. **Menssen et al. 2024:** Prediction intervals for overdispersed Poisson data and their application in medical and pre-clinical quality control arXiv:2404.05282 (under review). **Menssen and Rathjens 2024:** Prediction intervals for overdispersed binomial endpoints and their application to historical control data arXiv:2407.13296 (under review).

Take home messages

- ▶ Methodology in line with OECD and EFSA requirements
- ▶ Methodology is validated
- ▶ Applicable for other purposes
 - ▶ ADA cut-points
 - ▶ Sheward control charts

¹ **Menssen and Schaarschmidt 2019:** Prediction intervals for overdispersed binomial data with application to historical controls. *Statistics in Medicine*, 38(14), 2652-2663. **Menssen and Schaarschmidt 2022:** Prediction intervals for all of M future observations based on linear random effects models, *Statistica Neerlandica*, 76(3), 283-308. **Menssen et al. 2024:** Prediction intervals for overdispersed Poisson data and their application in medical and pre-clinical quality control arXiv:2404.05282 (under review). **Menssen and Rathjens 2024:** Prediction intervals for overdispersed binomial endpoints and their application to historical control data arXiv:2407.13296 (under review).

Thank you!

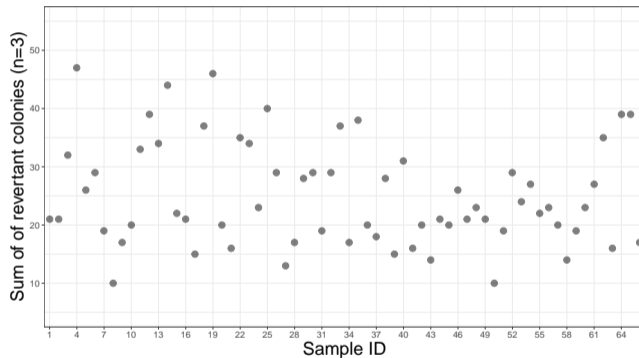
Appendix

Count data: No. of revertant colonies
(Ames test)

HCD

##	colonies	plates
## 1	27	3
## 2	14	3
## 3	17	3
## 4	29	3
## 5	21	3
## 6	17	3

Historical control data: TA 1537



¹Tarone 1982: The Use of Historical Control Information in Testing for a Trend in Poisson Means. Biometrics, 457-462.

Current trial

##	dose	colonies	plates
## 1	0.0	10	3
## 2	0.3	18	3
## 3	1.0	21	3
## 4	3.3	16	3
## 5	10.0	35	3

Aim

- ▶ Validation of the control group
- ▶ $y^* = 10$ revertant colonies
- ▶ $n^* = 3$ petri dishes

¹Tarone 1982: The Use of Historical Control Information in Testing for a Trend in Poisson Means. Biometrics, 457-462.

Quasi-Poisson assumption

- ▶ $\text{var}(Y_h) = \phi n_h \lambda$
- ▶ $\text{var}(Y^*) = \phi n^* \lambda$
- ▶ Overdispersion $\phi > 1$

Prediction standard error

- ▶ $\widehat{\text{se}}(\hat{y}^* - Y^*) = \sqrt{\widehat{\text{var}}(\hat{y}^*) + \widehat{\text{var}}(Y^*)} = \sqrt{\frac{\hat{\phi} n^{*2} \hat{\lambda}}{\sum_h n_h} + \hat{\phi} n^* \hat{\lambda}}$

¹Menssen et al. 2024: Prediction intervals for overdispersed Poisson data and their application in medical and pre-clinical quality control
arXiv:2404.05282 (under review).

BS-calibrated interval

$$l = n^* \hat{\lambda} - q_l \sqrt{\frac{\hat{\phi} n^{*2} \hat{\lambda}}{\sum_h n_h} + \hat{\phi} n^* \hat{\lambda}},$$

$$u = n^* \hat{\lambda} + q_u \sqrt{\frac{\hat{\phi} n^{*2} \hat{\lambda}}{\sum_h n_h} + \hat{\phi} n^* \hat{\lambda}}$$

¹Menssen et al. 2024: Prediction intervals for overdispersed Poisson data and their application in medical and pre-clinical quality control
arXiv:2404.05282 (under review).


```
qp_predint <- quasi_pois_pi(histdat = tarone_hcd,  
                             newdat = tarone_ad[1, 2:3])
```

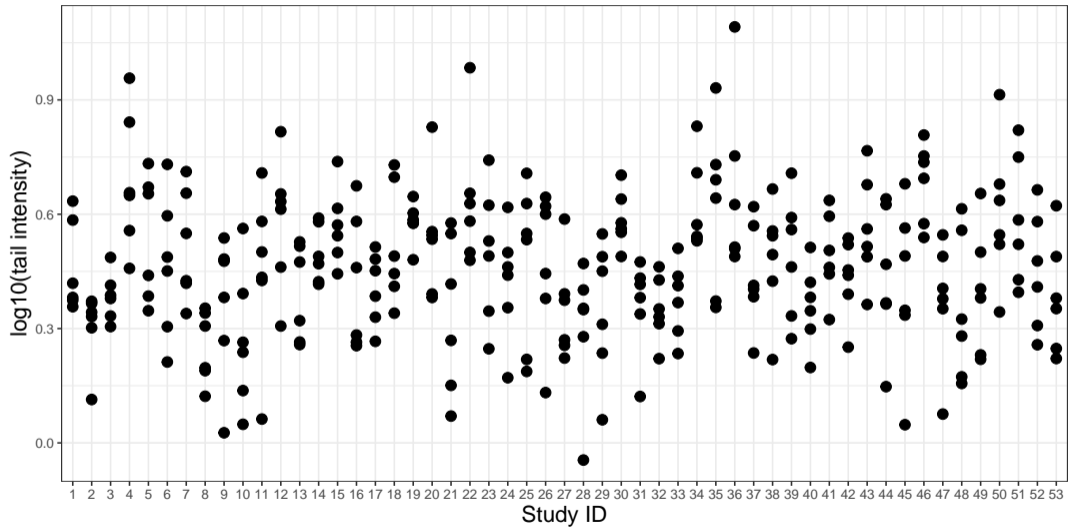
```
qp_predint$phi # 3.182295
```

```
summary(qp_predint)
```

```
## Pointwise 95 % prediction interval for one future observation
##
## colonies plates lower upper y_star_hat ql
## 1 10 3 10.04798 44.81337 25.06061 1.668496
## qu pred_se cover
## 1 2.195312 8.997699 FALSE
##
## ATTENTION: Not all future observations are covered
##
## Bootstrap calibration was done for each prediction limit separately
## using a modified version of Menssen and Schaarschmidt 2022
```

Continuous data: Comet assay

HCD: Comet assay



Experimental design

- ▶ \log_{10} (Tail intensity)
- ▶ Six rats per control group
- ▶ Rats nested in historical control groups

Model

- ▶ One way random effects model
- ▶ $y_{i(h)} = \mu + a_h + e_{i(h)}$
- ▶ $a_h \sim N(0, \sigma_a^2)$ σ_a^2 : Between study variation
- ▶ $e_{i(h)} \sim N(0, \sigma_e^2)$ σ_e^2 : Within study variation

Historical observations

$$\mathbf{Y} = \mathbf{1}\mu + \mathbf{ZU} + \epsilon$$

$$\mathbf{ZU} = (\mathbf{Z}_1, \dots, \mathbf{Z}_C) \begin{pmatrix} \mathbf{U}_1 \\ \vdots \\ \mathbf{U}_C \end{pmatrix} = \sum_{c=1}^C \mathbf{Z}_c \mathbf{U}_c$$

$$\text{var}(\mathbf{Y}) = \mathbf{\Sigma} = \sum_{c=1}^C \mathbf{Z}_c \mathbf{Z}_c^T \sigma_c^2 + \mathbf{I} \sigma_{C+1}^2$$

$$\mathbf{Y} \sim \text{MVN}(\mathbf{1}\mu, \mathbf{\Sigma})$$

$$Y \sim N(\mu, \sigma_1^2 + \sigma_2^2, \dots, \sigma_{C+1}^2)$$

¹Searle et al. 2006

Current observations

$$\mathbf{Y}^* = \mathbf{1}\mu + \mathbf{Z}^* \mathbf{U}^* + \boldsymbol{\epsilon}^*$$

$$\mathbf{Z}^* \mathbf{U}^* = (\mathbf{Z}^*_1, \dots, \mathbf{Z}^*_C) \begin{pmatrix} \mathbf{U}^*_1 \\ \vdots \\ \mathbf{U}^*_C \end{pmatrix} = \sum_{c=1}^C \mathbf{Z}^*_c \mathbf{U}^*_c$$

$$\text{var}(\mathbf{Y}^*) = \boldsymbol{\Sigma}^* = \sum_{c=1}^C \mathbf{Z}^*_c \mathbf{Z}^{*T}_c \sigma_c^2 + \mathbf{I}^* \sigma_{C+1}^2$$

$$\mathbf{Y}^* \sim \text{MVN}(\mathbf{1}\mu, \boldsymbol{\Sigma}^*)$$

$$Y^* \sim N(\mu, \sigma_1^2 + \sigma_2^2, \dots, \sigma_{C+1}^2)$$

Overview about the data

```
## 'data.frame':   318 obs. of  3 variables:  
## $ y_log10: num  0.357 0.374 0.634 0.585 0.419 ...  
## $ run    : Factor w/ 53 levels "1","2","3","4",...: 1 1 1 1 1 1 2 2 2 2  
## $ id     : Factor w/ 53 levels "6","12","18",...: 1 2 3 4 5 6 7 8 9 10
```


Model

- ▶ One way random effects model
- ▶ $y_{i(h)} = \mu + a_h + e_{i(h)}$

Fit the model

```
# Fit the model  
comet_fit <- lme4::lmer(y_log10 ~ 1 + (1|run), data=comet_hcd)
```

Prediction intervals with `lmer_pi_futmat()`

- ▶ Works with random effects models fit with `lme4::lmer()`
- ▶ Factorial designs (hierarchical, cross-classified, ...)
- ▶ Random effects must be specified as `(1|ran_eff)`
- ▶ Provides PI on observation level

Pointwise PI: One future rat

```
# Pointwise PI on log10 scale
pred_int_pw <- lmer_pi_futmat(model=comet_fit,
                             newdat=1,
                             traceplot=FALSE)
summary(pred_int_pw)
```

```
## Pointwise 95 % prediction interval for one future observation
```

```
##
```

```
##      lower      upper y_star_hat      q      pred_se
```

```
## 1 0.1081695 0.8180777 0.4631236 1.961172 0.1809908
```

```
##
```

```
## Bootstrap calibration was done following Menssen and Schaarschmidt 2022
```

Simultaneous PI: One concurrent control with six rats

- ▶ $\text{var}(\mathbf{Y}^*) = \boldsymbol{\Sigma}^* = \sum_{c=1}^C \mathbf{Z}_c^* \mathbf{Z}_c^{*T} \sigma_c^2 + \mathbf{I}^* \sigma_C^2 + 1$
- ▶ Provide list of \mathbf{Z}_c^* via `futmat_list`

Simultaneous PI: One concurrent control with six rats

```
# Variance components  
lme4::VarCorr(comet_fit)
```

```
## Groups   Name          Std.Dev.  
## run      (Intercept) 0.088005  
## Residual                0.157444
```

```
fml[["run"]]
```

```
##      [,1]  
## [1,]    1  
## [2,]    1  
## [3,]    1  
## [4,]    1  
## [5,]    1  
## [6,]    1
```

```
fml[["Residual"]]
```

```
##      [,1] [,2] [,3] [,4] [,5] [,6]  
## [1,]    1    0    0    0    0    0  
## [2,]    0    1    0    0    0    0  
## [3,]    0    0    1    0    0    0  
## [4,]    0    0    0    1    0    0  
## [5,]    0    0    0    0    1    0  
## [6,]    0    0    0    0    0    1
```

Simultaneous PI: One concurrent control with six rats

```
# Simultaneous PI on log10 scale  
pred_int_si1 <- lmer_pi_futmat(model=comet_fit,  
                               futmat_list=fml,  
                               traceplot=FALSE)
```

Simultaneous PI: One concurrent control with six rats

```
summary(pred_int_si1)
```

```
## Simultaneous 95 % prediction interval for 6 future observations
##
##           lower      upper y_star_hat      q      pred_se
## 1 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
## 2 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
## 3 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
## 4 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
## 5 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
## 6 -0.01543098 0.9416782  0.4631236 2.644082 0.1809908
##
## Bootstrap calibration was done following Menssen and Schaarschmidt 2022
```


Simultaneous PI: One concurrent control with six rats

```
# ggplot based graphic  
plot(pred_int_si1)+  
  ylab("log10(tail intensity) ") +  
  theme(text = element_text(size = 15))
```

Simultaneous 95 % prediction interval for 6 future observations

